

An Econometric Evaluation of the Stock Market Inter-linkages of BRICS Countries

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Abstract

In this paper an attempt has been made to examine the causal relationship between the returns of stock exchanges of BRICS countries and looks at the integration of the Indian stock market with the other BRICS nations. An Econometric analysis of the daily returns of the stock markets of BRICS countries for a period beginning from June 1, 2009 through March 31, 2015 is conducted in EViews. Line Charts and Unit-Root Test are applied to check the stationarity of the series; Granger's Causality Model and Johansen Co-integration Test are performed to find out the linkages between the markets under study. The results show that the returns at Indian stock exchange Granger Cause the return at Russia, Brazil and South Africa stock Exchanges. Neither the China Stock returns are Granger caused by India Stock Exchange nor is the Indian Exchange is Granger Caused by China Stock Exchange. The co movement of the stock market indices in different countries posed a big challenge as they reduce the benefit of diversification of investment. The Co-integration Test shows that the BRICS stock market is not co-integrated.

Key Words: Inter-linkages, Stationary, Granger's Causality, Co-integration

INTRODUCTION

After the practice of international portfolio diversification being established, institutional as well as individual investors from US and other developed and emerging

countries have become increasingly interested in foreign equity markets. However, in the present situation, global markets have tended to become more integrated as a result of common practice towards liberalization and deregulation in the capital markets of both developed as well as emerging countries. These basic changes lead to increase the correlations among the stock prices of various national markets all over the world that would ensure the reduction in the potential benefits from international portfolio diversification.

In the recent decade, markets, businesses, regions, and continents have become more interdependent upon one another. This phenomenon encourages wide range of financial services and raising funds throughout the world. The globalization of economic activity, the increased world wealth, and the reduction in transaction costs associated with the information revolution all direct investors to consider the newly emerging financial markets. This process has led to the introduction of public share offerings to nearly two dozen countries and spawned a global market culture among millions of new investors.

The World Bank (1997) argues that the world's financial markets are rapidly integrating into a single global marketplace as investors are driven to developing countries in the search for higher returns. However, history has shown that few once-emerging economies, such as the United States and Japan have been successful. This has served as motivation for investors to allocate their investment in emerging market

in order to enhance portfolio return and reducing risk through diversification. The risk and return in investing in emerging markets are significantly linked to the ability of these markets to develop economically.

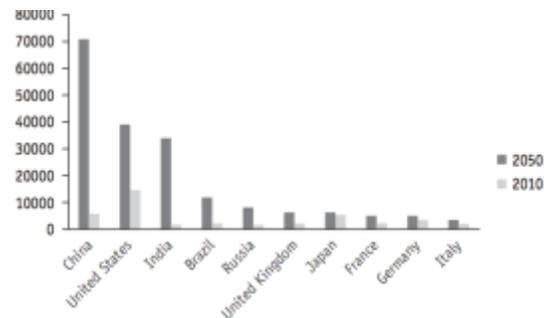
At the end of the century more countries than ever participate in capital markets. Moreover, companies all around the globe increasingly rely on the stock market to raise funds. This process is aided by the progression of countries to privatize their holdings and to transfer ownership from the state to private investors.

This phenomenon is not restricted to the United States, where the number of listed companies has increased by more than five times since 1990 to more than 20,000, or Western Europe, where governments auctioned off large portions of state-owned enterprises to the public. Estimates by Morgan Stanley Capital International reveal that the combined market capitalization for the United Kingdom, Germany and France increased by 250% in the last decade. Moreover, whereas ten years ago, China, the Soviet Union and its former Eastern Block of satellite countries have just abandoned command economies, and were lacking any stock market, at the end of 2008 exchanges in these countries enlisted more than 1,500 publicly traded companies.

About BRICS

BRICS is the acronym for an association of five major emerging national economies : Brazil, Russia, India, China and South Africa. The grouping was originally known as "BRIC" before the inclusion of South Africa in 2010. The acronym "BRIC" was coined by Jim O'Neill, a global economist working for Goldman Sachs, in a 2001 paper entitled "Building Better Global Economic BRICs". The acronym has come

into widespread use as a symbol of the shift in global economic power away from the developed G7 economies towards the developing world. The BRICS members are all developing countries, but they are distinguished by their large, fast-growing economies and have significant influence on regional and global affairs.



GDP of Major Powers 2010 vs 2050 (in billions of Dollars)

Source: <http://www.theatlantic.com/business/archive/2012/02/the-world-in-2050>

OBJECTIVES OF THE STUDY

The study aims to achieve the following objectives:

1. To study the return patterns in the equity markets of 'BRICS' Brazil, Russia, India, China and South Africa
2. To find out the linkages between the stock exchanges under study; and
3. To observe whether there exist enough opportunities for diversification among the stock exchanges of Brazil, Russia, India, China and South Africa.

LITERATURE REVIEW

The presence of co-movements among national stock markets limits the benefit of

international diversification. A number of studies have examined co-movements of international stock markets. A considerable amount of work has been done on the interrelationships among the world equity markets, especially focus on the major developed markets like United States and Japan. The performances of developed markets draw the world attention before and after the crash of 1987. The predominant feature of the Crash in 1987 was its global scale. The equity markets of the world reacted to the collapse of the Dow Jones index of the New York Stock Exchange with their own version of a crash. The Crash of October 1987 has made people realize that various national equity markets are so integrated. The developed markets e.g., the US exert a strong influence on other smaller markets. Eunand Shim (1989) use vector auto regression to study the interdependence among world equity markets and found evidence of co-movements among these markets with the US. By using a single equation model, Cheung and Mak (1992) examine the causal relationships between the Asian markets and the developed markets. They also found that the US market is an important global factor.

Asian capital markets are new stars among the emerging markets; many studies have been done in the 1990s and thereafter to study the co-movements between Asian markets and the stock markets in developed countries. Chan et al. (1992) conduct unit root and pairwise cointegration tests to examine the relationship among the Asian-Pacific markets and conclude that these Asian markets are not cointegrated. Chowdhury (1994) studies the relationships among four Newly Industrialized Economies (NIEs), Japan and US. He found that US market leads the four NIEs (Hong Kong, S. Korea, Singapore, and Taiwan) and there is significant link between the stock markets of Hong Kong and Singapore, Japan and US.

Phyletic (1995) found that there has been an increase integration of capital markets in Pacific Basin Region with US and Japan. Kwan, Sim and Cotsomitis (1995) studied the stock markets of Australia, Hong Kong, Japan, Singapore, South Korea, Taiwan, United Kingdom, United States and Germany and suggest that these markets are not weak form efficient as they found significant lead-lag relationships between these equity markets. Cashin et al. (1995) use the cointegration tests to assess the extent to which equity prices move similarly across countries and regions. They report increased integration of emerging equity markets since the beginning of 1990 via greater regionalization of national stock markets. Besides, if national stock markets are subject to a global shock that causes them to deviate from their long-run equilibrium relationship, it takes several months for the long-run relationship to reassert itself.

Chaudhuri (1997) investigates the common trends in seven Asian markets by using the Johansen cointegration methodology and reported a single trend. Palac-McMiken (1997) uses the monthly ASEAN market indices (Indonesia, Malaysia, Philippines, Singapore and Thailand) between 1987 and 1995 and found that with the exception of Indonesia, all the markets are linked with each other and that these markets are not collectively efficient. He suggested that there is still room for diversification across these markets despite evidence of interdependence among ASEAN stock markets. Masih and Masih (1999) found the high level of interdependence among markets in Thailand, Malaysia, US, Japan, Hong Kong, and Singapore from 1992 to 1997. Johnson and Soenen (2002) studied the equity market integration between the Japanese stock market and the other twelve equity markets in Asia. They found that the equity markets of Australia, Hong Kong, Malaysia, New

Zealand, and Singapore are highly integrated with the stock market in Japan. They also found the evidence that these Asian markets are becoming more integrated over time and that a higher import share as well as a greater differential in inflation rates, real interest rates, and GDP growth rates have negative effects on stock market co-movements between country pairs. The results from Yang et al. (2003) show that both long-run cointegration relationships and short-run causal linkages among US, Japan and ten Asian emerging markets have been more integrated after the Asian Financial Crisis.

The linkage among BRIC stock markets is a subject that has attracted wide attention.

Neil Manning

(Common trends and convergence? South East Asian equity markets, 1988–1999) using the VAR module had concluded his study as these circumstances, the Johansen approach says little about the dynamics of convergence although it clearly signals the interdependence of these markets.

Aamir R. Hashmi Interlinkages among South East Asian Stock Markets (A Comparison Between Pre- and Post-1997-Crisis Periods) using the data from 1 Jan 1990 to 31 July 1997 and apply the tool of the correlation test Multivariate Granger Causality Tests In order to have some idea about the properties of the data is compute

basic statistics for all seven series of returns for both pre- and post-crisis periods. Mean and median are generally lower for post crisis period. Variance is higher after the crisis for all seven markets. Noticeable are the post-crisis variance numbers (which are all greater than 4) for

KL, BK and JK. Changes in kurtosis are mixed and in four of these seven markets it has increased while for the remaining three it has decreased. Although the plots of return series (not shown) suggest stationary, the

test for unit roots and are able to reject the null of unit roots in all seven series of returns for both periods. Interlinkages among Asian markets have generally increased after the emergence of Asian financial crisis in 1997. Across the Asian Equity Markets (January 4, 1988 to December 31, 1999) using unit root test and co-integration test, concluded that for the level of the raw (un-trended) series, none of the tests, with the exception of India, reject the null hypothesis of non-stationarity at the 5 percent level. However, once the trend component is incorporated, the test indicates that all equity market indices, including the Indian index, are non-stationary in the levels. Consequently, co-integration tests are estimated in a manner that accommodates both intercept and trend terms. For the overall sample period, evidence from the max and Trace test statistic indicates that there are two significant cointegrating vectors for both Group I and Group II countries. An examination of the individual subperiod results (specifically, the max test) suggest that there is more than one cointegrating relationship with the inclusion of Japan and the United States in the system.

It appears that previous empirical studies on the relationship between world stock markets do not provide consistent results. The reasons for the inconsistent results are numerous, including the choice of markets, different sample periods, different frequency of observations, and the different methodologies employed.

RESEARCH METHODOLOGY

The present research is a study of the linkages between the stock exchanges of Brazil, Russia, India, China and South Africa. We use one stock exchange from each of the three countries as a representative of the respective country. The Brazil Stock Exchange

istakenasthe benchmarkofBraziland the index ofBM&FBOVESPAisused. ForRussia,theRussiaStockExchange(RTS)is takenasthe representativeexchangeand theAll Share PriceIndex(ASPI) is used. National Stock Exchange fromIndiaistakenasbenchmarkindexandthe Nifty GeneralIndexisused.ShanghaiStock Exchange(SSE)istakenasthebenchmarkstock exchangeforChina. The FTSE JSE is taken as a representative exchange for South Africa. Thedailyclosinglevelsofthefive representativendicesforaperiodbeginningfrom1 June 2009 through31 March2015 isconsidered the reference period.Inthisway,data of total70 monthsare takenfor thepurposeofthestudy. For the days when either of the stock market is closed, taking Indian Stock Market as base, the missing values are calculated by taking the average of the two nearest values. Datahavebeenanalyzedusingeconometrictools. Theanalysisofeconometricscanbeperformedonseriesofstationary nature.Inordertocheck whether or not the series are stationary, we prepare the line graph for each of the series. In order to further confirm the (stationary) nature of the series, correlogram is prepared for each of the series. Further, we perform the Augmented Dickey-Fuller test under the unit root test to finally confirm whether or not the series are stationary.Forthebasicunderstanding ofUnitroot testing, wemaylookat the following equation

$$y_t = \rho y_{t-1} + x_t' \delta + \varepsilon_t, (1.1)$$

Where x_t are optional exogenous regressors which may consist of constant, or a constant and trend, and δ are parameters to be estimated, and the ε_t are assumed to be white noise. If $|\rho|$ is a nonstationarity series and the variance of y_t increases with time and approaches infinity. If

$|\rho| < 1$, y_t is a trend-stationary series. Thus, we evaluate the hypothesis of trend-stationary by testing whether the absolute value of $|\rho|$ is strictly less than one. The Standard Dickey-Fuller test is carried out by estimating equation (1.2) after subtracting y_{t-1} from both sides of the equation.

$$\Delta y_t = \alpha y_{t-1} + x_t' \delta + \varepsilon_t, (1.2)$$

Where $\alpha = \rho - 1$.

The null and alternative hypotheses may be written as,

$$H_0: \alpha = 0 \quad H_1: \alpha < 0 (1.3)$$

In order to make the series stationary, we take the log of the five series and arrive at the daily return of the series. All the remaining analysis is performed at the daily return (log of the series) of the five exchanges. These variables are named as LOGBZ, LOGCH, LOGRUS, LOGIN and LOGSA.

At the stationary log series of the five stock exchanges, we perform the Granger's causality model in order to observe whether the return at each stock exchange Granger causes the return at the stock exchanges. The Granger (1969) approach to the question of whether x causes y is to see how much of the current y can be explained by past values of y and then to see whether adding lagged values of x can improve the explanation. y is said to be Granger-caused by x if x helps in the prediction of y , or equivalently if the coefficients on the lagged x 's are statistically significant. It is pertinent to note that two-way causation is frequently the case; x Granger causes y and y Granger causes x . It

is important to note that the statement " x Granger causes y " does not imply that y is the effect or the result of x . Granger causality measures precedence and information content but does not by itself indicate causality in the more common use of the term. In Granger's Causality, there are bivariate regressions of the under-mentioned form—

$$y_t = \alpha_0 + \alpha_1 y_{t-1} + \dots + \alpha_k y_{t-k} + \beta_1 x_t -$$

$$1 + \dots + \beta_1 x_{t-1} + \varepsilon_t$$

$$x_t = \alpha_0 + \alpha_1 x_{t-1} + \dots + \alpha_l x_{t-l} + \beta_1 y_{t-1} + \dots + \beta_l y_{t-l} + \mu_t \quad (1.4)$$

for all possible pairs of (x, y) series in the group. In equation (1.4), we take lags ranging from 1 to l. In Granger's model, one can pick a lag length, l, that corresponds to reasonable beliefs about the longest time over which one of the variables could help predict the other.

The reported F-statistics are the Wald statistics for the joint hypothesis:

$$\beta_1 = \beta_2 = \dots = \beta_l = 0 \quad (1.5)$$

for each equation.

The null hypothesis is that x does not Granger-cause y in the first regression and that y does not Granger-cause x in the second regression.

Data Analysis

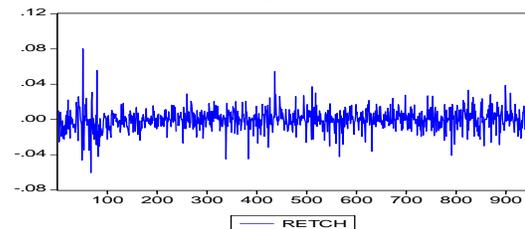
The data taken for the returns of the stock indices of the BRICS, ie, Brazil, Russia, India, China and South Africa is statistically evaluated in this section. The Descriptive Statistics for the return series of the countries under study is shown in Table 1. The descriptive statistics shows that the return indices are not normally distributed. These are positively skewed which means these series have a long right tail. The kurtosis for the normal distributed series is 3. The Kurtosis is more than 3 for Russia and China, it infers that these return series are peaked (leptokurtic) and for India, China and South Africa it can be inferred that these return series are Mesokurtic as their Kurtosis value is less than 3.

Table 1

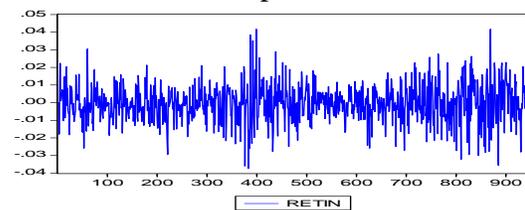
Descriptive Statistics						
	Total	India	South Africa	Brazil	Russia	China
Mean	734984.7	6198.471	36544.16	55251.71	1468.471	2337.397
Median	734982.5	5873.225	35950.5	55082.84	1462.82	2240.51
Maximum	735487	8996.25	47325	68394.33	1838.18	3786.57
Minimum	734288	4544.2	25193	44965.66	1237.43	1950.01
Std. Dev.	404.3378	1155.218	6527.813	4833.992	101.8531	344.159
Skewness	0.012828	0.938907	0.128116	0.207558	0.855373	1.842381
Kurtosis	1.810875	2.671198	1.611822	2.541951	4.890889	6.252295
Jarque-Bera	55.29028	142.0403	77.8811	14.9348	160.8788	944.0536
Probability	0	0	0	0.004571	0	0
Sum	6.80E+08	5814165	34278419	51826100	1377426	2192478
Observations	938	938	938	938	938	938

The five series representing the stock indices of the BRICS are statistically analysed in Eviews. For Econometric analysis, it is essential to make sure that the series under study are stationary. In order to make the series stationary, log is taken. All the analysis has been done on the log series. The series is stationary at first difference. In order to make the series stationary, line graphs are prepared. The line graphs are shown below

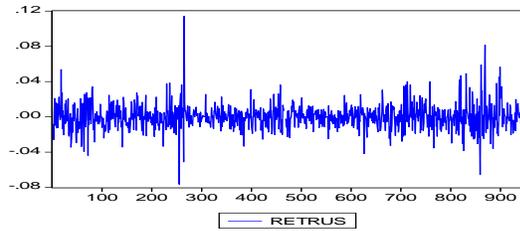
Graph 1



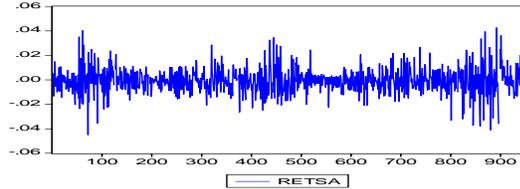
Graph 2



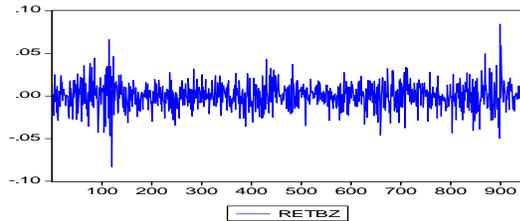
Graph 3



Graph 4



Graph 5



Figures 1 to 5 demonstrate the line graphs of the returns of the five stock exchanges. RETCH, RETIN, RETRUS, RETSA, RETBZ represents the first difference series. It is indicated from the figures that returns at all the five stock exchanges are stationary. In order to further check the stationarity of the three series, we perform the correlogram and the Unit Root Test in order to further confirm the same. Table 2-6 summarizes the unit root test of return.

Augmented Dickey Fuller Unit Root Test

Table 2

Augmented Dickey-Fuller Test Equation				
Dependent Variable: D(LOGBZ,2)				
Method: Least Squares				
Sample (adjusted): 3 946				
Included observations: 944 after adjustments				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(LOGBZ(-1))	-1.051282	0.032551	-32.29628	0
C	0.000238	0.000493	0.482676	0.6294
R-squared	0.525453	Mean dependent var	-1.53E-05	
Adjusted R-squared	0.524949	S.D. dependent var	0.021983	
S.E. of regression	0.015151	Akaike info criterion	-5.53932	
Sum squared resid	0.216252	Schwarz criterion	-5.52905	
Log likelihood	2616.561	F-statistic	1043.05	
Durbin-Watson stat	1.994581	Prob(F-statistic)	0	

Table 3

Augmented Dickey-Fuller Test Equation				
Dependent Variable: D(LOGIN,2)				
Method: Least Squares				
Sample (adjusted): 3 946				
Included observations: 944 after adjustments				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(LOGIN(-1))	-0.904316	0.032442	-27.8748	0
C	-0.0004	0.000341	-1.17091	0.2419
R-squared	0.452008	Mean dependent var	7.76E-06	
Adjusted R-squared	0.451427	S.D. dependent var	0.014141	
S.E. of regression	0.010474	Akaike info criterion	-6.27779	
Sum squared resid	0.103335	Schwarz criterion	-6.26752	
Log likelihood	2965.118	F-statistic	777.0042	
Durbin-Watson stat	1.990779	Prob(F-statistic)	0	

Table 4

Augmented Dickey-Fuller Test Equation				
Dependent Variable: D(LOGRUS,2)				
Method: Least Squares				
Sample (adjusted): 3 946				
Included observations: 941 after adjustments				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(LOGRUS(-1))	-1.015238	0.03263	-31.1139	0
C	2.19E-05	0.000477	0.045798	0.9635
R-squared	0.507623	Mean dependent var	1.31E-05	
Adjusted R-squared	0.507098	S.D. dependent var	0.020847	
S.E. of regression	0.014636	Akaike info criterion	-5.60851	
Sum squared resid	0.201152	Schwarz criterion	-5.59821	
Log likelihood	2640.803	F-statistic	968.0746	
Durbin-Watson stat	2.000189	Prob(F-statistic)	0	

Table 5

Augmented Dickey-Fuller Test Equation				
Dependent Variable: D(LOGSA,2)				
Method: Least Squares				
Sample (adjusted): 4 946				
Included observations: 923 after adjustments				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(LOGSA(-1))	-1.153335	0.049118	-23.48079	0
D(LOGSA(-1),2)	0.049125	0.033036	1.487003	0.1374
C	-0.000665	0.000343	-1.938935	0.0523
R-squared	0.548661	Mean dependent var		-6.02E-05
Adjusted R-squared	0.54768	S.D. dependent var		0.015463
S.E. of regression	0.010399	Akaike info criterion		-6.29087
Sum squared resid	0.099498	Schwarz criterion		-6.27518
Log likelihood	2906.238	F-statistic		559.1905
Durbin-Watson stat	1.979944	Prob(F-statistic)		0

Table 6

Augmented Dickey-Fuller Test Equation				
Dependent Variable: D(LOGCH,2)				
Method: Least Squares				
Sample (adjusted): 20 946				
Included observations: 887 after adjustments				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
D(LOGCH(-1))	-0.759162	0.130527	-5.816124	0
D(LOGCH(-1),2)	-0.253254	0.127826	-1.981241	0.0479
C	-5.10E-05	0.000405	-0.125979	0.8998
R-squared	0.524125	Mean dependent var		-1.48E-05
Adjusted R-squared	0.514256	S.D. dependent var		0.017215
S.E. of regression	0.011998	Akaike info criterion		-5.98695
Sum squared resid	0.124953	Schwarz criterion		-5.88439
Log likelihood	2674.21	F-statistic		53.11151
Durbin-Watson stat	1.999118	Prob(F-statistic)		0

With the Unit-Root Test, the null hypothesis that series Return at the benchmark exchanges and indices has a unit-root is tested. Probability value of less than 0.05 in table implies that the Null hypothesis is rejected and the variable does not have a unit-root, which confirms that the series is stationary. Similar kind of results is visible from the Augmented Dickey-Fuller test in table. After confirming that the five series are stationary in nature, the pairwise Granger's Causality analysis on the benchmark indices is performed. Table 7 presents the findings of pairwise Granger's Causality for the stock exchanges under study.

Table 7

Pairwise Granger Causality Tests

Null Hypothesis:	Obs	F-Statistic	Probability	Causal Relationship
RETRUS does not Granger Cause RETIN	939	0.61277	0.54207	No
RETIN does not Granger Cause RETRUS		8.72662	0.00018	Yes
RETSA does not Granger Cause RETIN	923	0.75633	0.46968	No
RETIN does not Granger Cause RETSA		10.3478	3.60E-05	Yes
RETC does not Granger Cause RETIN	935	0.60293	0.54742	No
RETIN does not Granger Cause RETC		0.33782	0.71341	No
RETBZ does not Granger Cause RETIN	943	0.26807	0.76492	No
RETIN does not Granger Cause RETBZ		27.0658	3.70E-12	Yes

Table 7 presents the findings of the Granger's Causality model to the benchmark indices from Brazil, Russia, India, China and South Africa. Null hypothesis in the case of Granger's causality model is that 'A' does not Granger cause 'B'. On those lines, table tests the hypotheses about the four stock exchanges in pair with India. In table we particularly observe the probability value (shown in the last column). We test the null hypothesis at 5% level of significance for which the probability value of less than 0.05 implies that the null hypothesis can be rejected while for the probability value of more than 0.05, we shall accept the null hypothesis. The results show that the probability value for the hypothesis at 2 lags.

The Probability value of the hypothesis: RETRUS does not Granger Cause RETIN, RETSA does not Granger Cause RETIN, RETC does not Granger Cause RETIN, RETIN does not Granger Cause RETC, RETBZ does not Granger Cause RETIN is more than 0.05, which implies that null hypothesis cannot be rejected. Whereas the probability value of hypothesis: RETIN does not Granger Cause RETRUS, RETIN does not Granger Cause RETSA, RETIN does not Granger Cause RETBZ is lower than 0.05, which means the null hypothesis

in can be rejected.

This implies that the return at Indian stock exchange Granger Cause the return at Russia, Brazil and South Africa stock Exchanges. Neither the China Stock returns are Granger caused by India Stock Exchange nor is the Indian Exchange is Granger Caused by China Stock Exchange.

Table 8
Result of Cointegration Test

JOHNSON COINTEGRATION TEST				
Sample (adjusted): 7946				
Included observations: 884 after adjustments				
Trend assumption: Linear deterministic trend				
Series: RETBZ RETIN RETSA RETCH RETRUS				
Lags interval (in first differences): 1 to 4				
Unrestricted Cointegration Rank Test (Trace)				
Hypothesized		Trace	0.05	
No. of CE(s)	Eigenvalue	Statistic	Critical Value	Prob.**
None*	0.21551	887.8303	63.81863	0.0001
At most 1*	0.190215	673.2643	47.85613	0.0001
At most 2*	0.173553	486.7533	29.79707	0.0001
At most 3*	0.165216	311.7982	15.49471	0.0001
At most 4*	0.15813	152.1631	3.841466	0
Trace test indicates 5 cointegrating eqn(s) at the 0.05 level				
* denotes rejection of the hypothesis at the 0.05 level				
**MacKinnon-Haug-Michelis (1993) p-values				
Unrestricted Cointegration Rank Test (Maximum Eigenvalue)				
Hypothesized		Max-Eigen	0.05	
No. of CE(s)	Eigenvalue	Statistic	Critical Value	Prob.**
None*	0.21551	214.5653	33.87687	0.0001
At most 1*	0.190215	186.5116	27.58434	0.0001
At most 2*	0.173553	174.3951	21.13162	0.0001
At most 3*	0.165216	153.6351	14.2646	0.0001
At most 4*	0.15813	152.1631	3.841466	0
Max-eigenvalue test indicates 5 cointegrating eqn(s) at the 0.05 level				
* denotes rejection of the hypothesis at the 0.05 level				
**MacKinnon-Haug-Michelis (1993) p-values				
Unrestricted Cointegrating Coefficients (normalized by b***1*b1):				
RETBZ	RETIN	RETSA	RETCH	RETRUS
-85.01158	-33.43208	260.8701	36.5393	42.73627
-128.3378	105.1046	50.86313	41.44346	-103.2636
-79.55115	-103.8634	-112.3474	31.77437	140.3138
43.05112	158.3331	-132.3852	115.4973	-14.41734
-79.88624	139.933	-65.77137	-112.7967	105.0769
Unrestricted Adjustment Coefficients (alpha):				
D(RETBZ)	D(RETIN)	D(RETSA)	D(RETCH)	D(RETRUS)
0.000321	0.005686	0.00234	-0.002174	0.001722
-0.001444	3.57E-05	0.002064	-0.003102	-0.00234
-0.004713	0.001531	0.001548	-8.74E-05	0.000281
-0.002232	-0.000663	-0.001674	-0.003853	0.002338
-0.003595	0.004473	-0.002223	-0.001288	-0.00267

Comparing the critical values with the statistics, it can be interpreted that the

BRICS nations Stock market is not co-integrated. If the statistics values are lesser than the critical values, it infers that there is a co-integration. In this case, all the statistics values are much higher than their respective critical values. It can be concluded that the BRICS stock markets are not co-integrated.

Conclusion

The Econometric analysis on the returns from the stock exchanges in Brazil, Russia, India, China and South Africa namely– BM&FBOVESPA Stock Exchange, RST Stock Exchange, National Stock Exchange, Shanghai Stock Exchange and Johannesburg stock Exchange (FTSE/JSE) is conducted. In order to check the stationarity of the five series, the correlogram and the Unit Root Test are performed. We test the null hypothesis at 5% level of significance for which the probability value of less than 0.05 implies that the null hypothesis can be rejected. The Probability value of the hypothesis: RETRUS does not Granger Cause RETIN, RETSA does not Granger Cause RETIN, RETCH does not Granger Cause RETIN, RETIN does not Granger Cause RETCH, RETBZ does not Granger Cause RETIN is more than 0.05, which implies that null hypothesis cannot be rejected. Whereas the probability value of hypothesis: RETIN does not Granger Cause RETRUS, RETIN does not Granger Cause RETSA, RETIN does not Granger Cause RETBZ is lower than 0.05, which means the null hypothesis in can be rejected.

This implies that the return at Indian stock exchange Granger Cause the return at Russia, Brazil and South Africa stock Exchanges. Neither the China Stock returns are Granger caused by India Stock Exchange nor is the Indian Exchange is

Granger Caused by China Stock Exchange. The result of the co-integration test inferred that the stock Markets of the BRICS countries are not co-integrated. The presence of co-movements among national stock markets limits the benefit of international diversification.

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